

(4.4)

c)

Power supply provides up to 40W

$$P_r(X > 40W) = \left(\frac{1}{4}\right)\left(\frac{1}{4}\right)\left(\frac{1}{4}\right)\left(\frac{1}{4}\right)\left(\frac{1}{4}\right) = \left(\frac{1}{4}\right)^5 = 0.00977$$

↓
total power required

↓
when all 5 beds draw power at the same time

(H) or (T)

(H) (H) (H) (H) (H) $\rightarrow \left(\frac{1}{2}\right)^5$

HW3 : 6.4 ; 7.1 ; 8.2 ; 9.3 (Ch.2) due 3/11

(2.5.1)

Gaussian Random voltage: $X : \begin{cases} \bar{X} = 5 & (\text{mean}) \\ \sigma_x^2 = 16 & (\text{variance}) \end{cases}$

$$a) P_r(X > 0) = 1 - P_r(X \leq 0)$$

$$= 1 - F(0) = 1 - \Phi\left(\frac{0 - \bar{X}}{\sigma_x}\right) = 1 - \left[1 - \Phi\left(\frac{\bar{X}}{\sigma_x}\right)\right] = \Phi\left(\frac{\bar{X}}{\sigma_x}\right) = \Phi\left(\frac{5}{4}\right) = \Phi(1.25) = 0.8944$$

(App D pg 432)

Property of Φ :

$$\boxed{\Phi(-y) = 1 - \Phi(y)} ; \Phi(y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^y e^{-\frac{u^2}{2}} du$$

$$\text{Proof: } \Phi(-y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{-y} e^{-\frac{u^2}{2}} du = \frac{1}{\sqrt{2\pi}} \left\{ \underbrace{\int_{-\infty}^{\infty} e^{-\frac{u^2}{2}} du}_{2 \int_0^{\infty} e^{-\frac{u^2}{2}} du} - \int_{-y}^{\infty} e^{-\frac{u^2}{2}} du \right\}$$

$$\frac{2 \Gamma(\frac{1}{2})}{2 \sqrt{\frac{1}{2}}} = \sqrt{2\pi}$$

$$= 1 - \frac{1}{\sqrt{2\pi}} \int_{-y}^{\infty} e^{-\frac{u^2}{2}} du = 1 - \frac{1}{\sqrt{2\pi}} \int_y^{\infty} e^{-\frac{u^2}{2}} (-du)$$

change of dummy $u \rightarrow -u$

$$= 1 - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^y e^{-\frac{u^2}{2}} du = \Phi(y)$$

Now, find $\Pr(\bar{X} > 0)$ using $Q(x) \rightarrow$ App E pg 434.

Recall: $1 = \Phi(y) + Q(y)$

$$\Pr(\bar{X} > 0) = 1 - F(0) = \Phi\left(\frac{\bar{X}}{\sigma_{\bar{X}}}\right) = \Phi(1.25) = 1 - Q(1.25) \\ = 1 - 0.1056 = 0.8944$$

App E
pg 434

Related Density Functions (to the Gaussian Density Function)

1) Power Distribution:

$$P = I^2 R$$

$\downarrow$ constant, fixed

$I =$ Gaussian Random Variable

$\rightarrow$ What is the Density Function for the power P ?

$$\text{Recall: } Y = X^2 \rightarrow f_Y(y) = \frac{1}{2\sqrt{y}} [f_X(\sqrt{y}) + f_X(-\sqrt{y})]$$

$$\text{Now: } f_I(i) = \frac{1}{\sqrt{2\pi}\sigma_I} e^{-\frac{(i-\bar{I})^2}{2\sigma_I^2}} \left\{ \begin{array}{l} \text{Mean: } \bar{I} \\ \text{Variance is } \sigma_I^2 \end{array} \right.$$

$$\rightarrow f_P\left(\frac{P}{R}\right) = \frac{1}{2\sqrt{\frac{P}{R}}} [f_I(\sqrt{\frac{P}{R}}) + f_I(-\sqrt{\frac{P}{R}})]$$

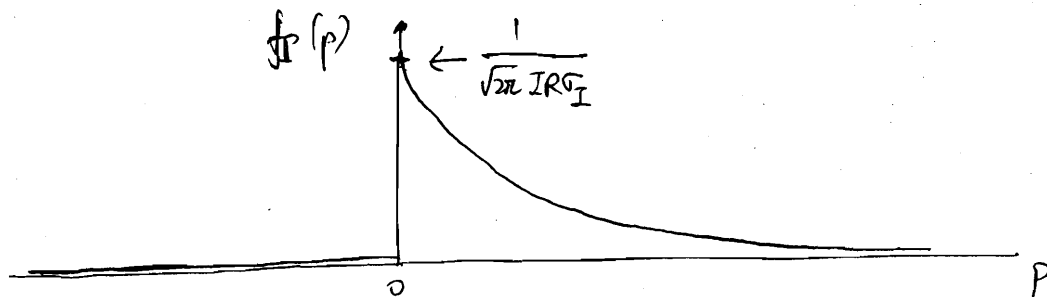
$$\rightarrow f_P(P) = \frac{1}{R} f\left(\frac{P}{R}\right) = \frac{1}{2\sqrt{PR}} [f_I(\sqrt{\frac{P}{R}}) + f_I(-\sqrt{\frac{P}{R}})]$$

linear connection:

$$Y = aX$$

$$\rightarrow \boxed{\text{Assume } \bar{I} = 0} \rightarrow f_P(p) = \frac{1}{2\sqrt{PR}} \left[2 \frac{1}{\sqrt{2\pi}\sigma_I} e^{-\frac{(\sqrt{\frac{P}{R}})^2}{2\sigma_I^2}} \right]$$

$$\rightarrow \text{Density function for the Power } \left\{ \begin{array}{l} f_P(p) = \frac{1}{\sqrt{\pi} IR \sigma_I} e^{-\frac{P}{2R\sigma_I^2}} : p \geq 0 \\ 0 : p < 0 \end{array} \right.$$



What are $\bar{P}$ (mean) & σ_P^2 (variance)?

Mean: $\bar{P} = \int_0^{\infty} dp \, p f_P(p) = \frac{1}{\sqrt{2\pi} R \sigma_I} \int_0^{\infty} dp \, \frac{p}{\sqrt{p}} e^{-\frac{p}{2R\sigma_I^2}} = \frac{1}{\sqrt{2\pi} R \sigma_I} \int_0^{\infty} dp \, p^{1/2} e^{-\frac{p}{2R\sigma_I^2}}$

$f_P(p) = 0$ when $p < 0$

$$\int_0^{\infty} dx \, x^n e^{-ax} = \frac{\Gamma(n+1)}{a^{n+1}} = \frac{n!}{a^{n+1}} \quad \text{if } n \text{ integer}$$

In our integral for $\bar{P}$ $\begin{cases} n = \frac{1}{2} \\ a = \frac{1}{2R\sigma_I^2} \end{cases} \rightarrow \bar{P} = \frac{1}{\sqrt{2\pi} R \sigma_I} \cdot \frac{\Gamma(\frac{1}{2}+1)}{(\frac{1}{2R\sigma_I^2})^{3/2}} = \frac{\frac{1}{2} \Gamma(\frac{1}{2})}{\sqrt{2\pi} R \sigma_I (\frac{1}{2R\sigma_I^2})^{3/2}}$

$\Gamma(\frac{1}{2}) = \sqrt{\pi}$

$\rightarrow \bar{P} = \frac{2R\sigma_I^2 \sqrt{\pi}}{\sqrt{\pi}} \rightarrow \boxed{\bar{P} = R\sigma_I^2}$ (we assumed $\bar{I} = 0$)

Makes sense?

$\bar{P} = R \overline{(I - \bar{I})^2} = R \overline{I^2} = \overline{RI^2}$
 $\downarrow$
 $R \text{ is constant.}$
 Yes, since $P = I^2 R \rightarrow \bar{P} = \overline{I^2 R}$

Variance

$\sigma_P^2 = \overline{(P - \bar{P})^2} = \overline{P^2} - \bar{P}^2 = \overline{P^2} - R^2 \sigma_I^4$

$\Gamma(\alpha+1) = \alpha \Gamma(\alpha)$
 property of Γ function

↑
 averaging algebra

Need to find now

use definition: $\overline{I^2} = \int_0^{\infty} dp \, p^2 f_P(p) = \frac{1}{\sqrt{2\pi} R \sigma_I} \int_0^{\infty} dp \, p^{3/2} e^{-\frac{p}{2R\sigma_I^2}} = \frac{1}{\sqrt{2\pi} R \sigma_I} \frac{\Gamma(\frac{3}{2}+1)}{(\frac{1}{2R\sigma_I^2})^{5/2}}$

$\Rightarrow \overline{P^2} = \frac{\frac{3}{2} \Gamma(\frac{3}{2})}{(\frac{1}{2R})^2 \sqrt{\pi} \frac{1}{\sigma_I^4}} = \frac{\frac{3}{2} \frac{1}{2} \Gamma(\frac{1}{2})}{\frac{\sqrt{\pi}}{(2R)^2 \sigma_I^4}} = \frac{\frac{3}{4} \sqrt{\pi} (2R)^2 \sigma_I^4}{\sqrt{\pi}} = \boxed{3R^2 \sigma_I^4}$

$$\rightarrow \sigma_P^2 = \overline{P^2} - \bar{P}^2 = 3R^2 \sigma_I^4 - R^2 \sigma_I^4 = \boxed{2R^2 \sigma_I^4}$$

Variance for Power P
when $\bar{I} = 0$

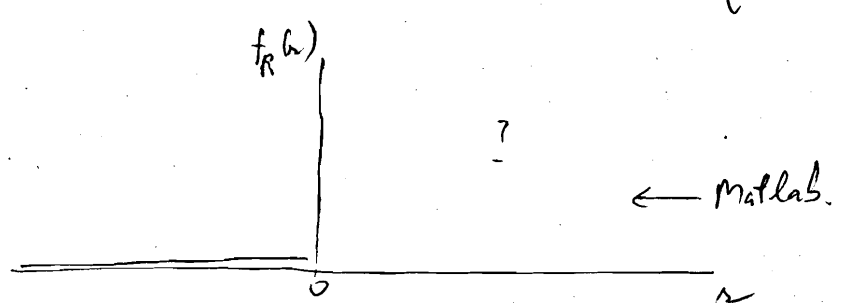
2) Rayleigh Distribution:

Hypothesis on length of a 2D vector (X, Y) , where X & Y are Gaussian random variables with zero means and same variance: $\sigma_x^2 = \sigma_y^2 = \sigma^2$, would follow the Rayleigh distribution, that means: $R = \sqrt{X^2 + Y^2}$, follows Rayleigh distribution:

Recall: $z = \sqrt{w}$, if W is Gaussian R. variable $\rightarrow f_z(z) = f_w(w) \left| \frac{dw}{dz} \right|$

$$\frac{dz}{dw} = -\frac{1}{2w^{1/2}} = -\frac{1}{2z} \rightarrow \left| \frac{dw}{dz} \right| = 2z \rightarrow f_z(z) = f_w(z^2) 2z$$

$$\rightarrow f_R(r) = f_{X^2+Y^2}(r^2) 2r = \begin{cases} \frac{1}{\sigma^2} e^{-\frac{r^2}{2\sigma^2}} & r \geq 0 \\ 0 & r < 0 \end{cases} \leftarrow \text{added for common sense}$$



Not symmetrical like the Gaussian dens. funct. for X & Y .

$$\bar{R} = \int_0^\infty dr \, r f_R(r) = \frac{1}{\sigma^2} \int_0^\infty dr \, r^2 e^{-\frac{r^2}{2\sigma^2}} = \frac{1}{\sigma^2} \frac{\Gamma(\frac{3}{2})}{2 \left(\frac{1}{2\sigma^2}\right)^{\frac{3}{2}}} \frac{\sigma \sqrt{\pi}}{2^{-\frac{1}{2}}}$$

Recall: $\int_0^\infty dx \, x^n e^{-x^2} = \frac{\Gamma(\frac{n+1}{2})}{2 \, 2^{\frac{n+1}{2}}}$

$$\rightarrow \boxed{\bar{R} = \sigma \sqrt{\frac{\pi}{2}}}$$

$$\begin{cases} n=2 \\ r^2 = \frac{1}{2\sigma^2} \end{cases}$$

$$\overline{R^2} = \int_0^\infty dr \, r^2 f_R(r) = \frac{1}{\sigma^2} \int_0^\infty dr \, r^3 e^{-\frac{r^2}{2\sigma^2}} = \frac{1}{\sigma^2} \frac{\Gamma(2)}{2\left(\frac{1}{2\sigma^2}\right)^{\frac{4}{2}}}$$

$\downarrow$
 $n=3$
 $r^2 = \frac{1}{2\sigma^2}$

$$= \frac{2}{\frac{\sigma^2}{\sigma^4} 2 \frac{2}{2^2}} = 2\sigma^2$$

$$\rightarrow \text{Variance} = \overline{R^2} - (\overline{R})^2 = 2\sigma^2 - \left(\sigma\sqrt{\frac{\pi}{2}}\right)^2$$

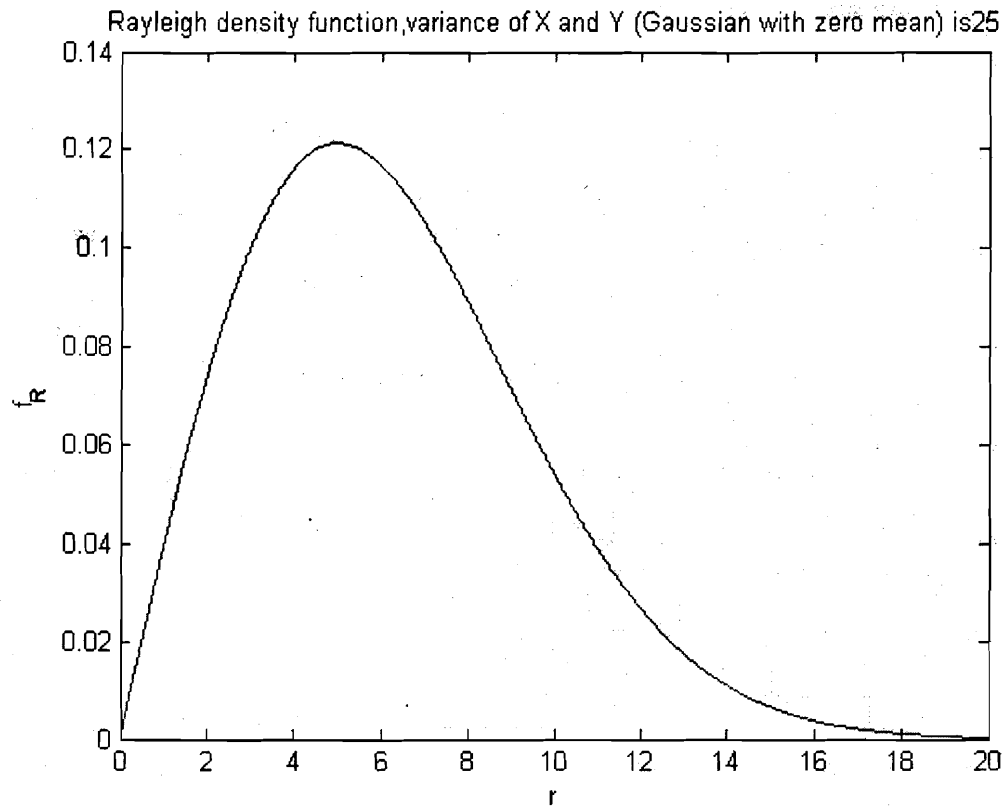
$$= \left(2 - \frac{\pi}{2}\right) \sigma^2 = 0.429 \sigma^2$$

$$\boxed{\sigma_R^2 = 0.429 \sigma^2}$$

Rayleigh

variance of Gaussian for $\bar{X}$ & $\bar{Y}$. (mean is 0)

Engin 322
March 6, 2008



%Plot of Rayleigh density function

```
ra=0:0.01:20;  
var=25;  
for i=1:2001  
    r=i/100;  
    fR(i)=(r)/var*exp(-r^2/(2*var));  
end  
figure(1), plot(ra,fR); title(strcat('Rayleigh density function, ',  
'variance of X and Y (Gaussian with zero mean) is', num2str(var)))  
xlabel('r'), ylabel('f_R')
```

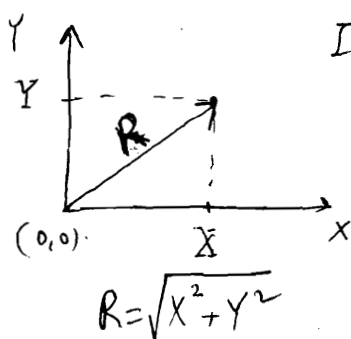
→ (6.4) → $X \rightarrow v_x ; Y \rightarrow v_y$ Gaussian Random Variables.
 zero mean ; $\sigma_x = \sigma_y = \sigma = 4 \text{ ft/s}$

a) Most probable speed $v = \sqrt{v_x^2 + v_y^2}$
 peak of the Rayleigh density function
 or max.

b) Mean value of speed: $\bar{v}$ (we derived formula
 for mean value of
 Rayleigh dens. function)

$$c) P_r(v > 10 \text{ ft/s}) = 1 - \underbrace{P_r(v \leq 10 \text{ ft/s})}_{F_v(10 \text{ ft/s})}$$

2D plane



If X, Y are Gaussian Random Variables with zero means and standard deviation $\sigma_x = \sigma_y = \sigma$

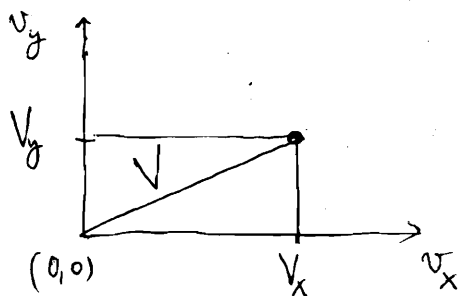
→ R is a random variable described by Rayleigh distribution: $f_R(r) = \frac{r}{\sigma^2} e^{-\frac{r^2}{2\sigma^2}} \quad r \geq 0$

Recall

$$\begin{cases} \bar{R} = \sigma \sqrt{\frac{\pi}{2}} \\ \sigma_R^2 = 0.429 \sigma^2 \end{cases}$$

Notation: $\begin{cases} \rightarrow \text{Capital letters} \rightarrow \text{Random variable} \\ \rightarrow \text{Lower-case letters} \rightarrow \text{values or argument of density functions or distribution functions} \end{cases}$

Another situation we can use Rayleigh distribution is problem (6.4)



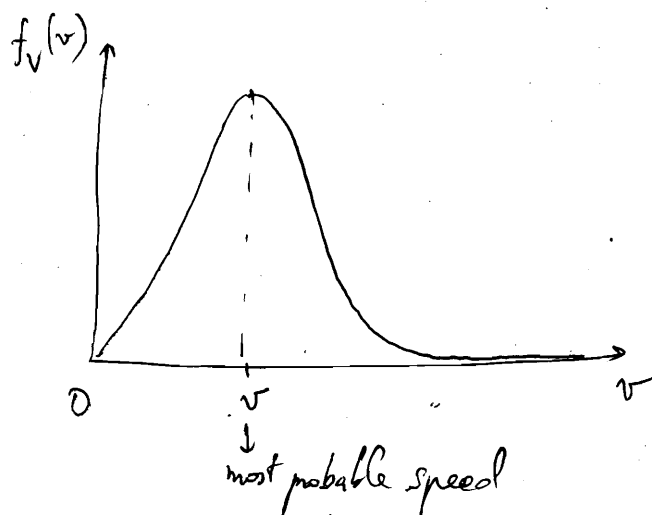
If V_x, V_y are Gaussian Random Variables with zero means and standard deviation $\sigma_{V_x} = \sigma_{V_y} = \sigma$

$\rightarrow V = \sqrt{V_x^2 + V_y^2}$ is the speed, is a random variable described by the Rayleigh distribution:

$$f_V(v) = \frac{v}{\sigma^2} e^{-\frac{v^2}{2\sigma^2}} \quad v \geq 0$$

$$\bar{V} = \sigma \sqrt{\frac{\pi}{2}}$$

$$\sigma_V^2 = 0.429 \sigma^2$$



(6.4)

a)

?

Since the Rayleigh dist. is NOT symmetric, the most probable value is NOT the mean value (which happens in the symmetric Gaussian distribution)

$$f_V(v) = \frac{v}{\sigma^2} e^{-\frac{v^2}{2\sigma^2}} \quad (\sigma \text{ is known: it's the standard deviation of } V_x \text{ and } V_y)$$

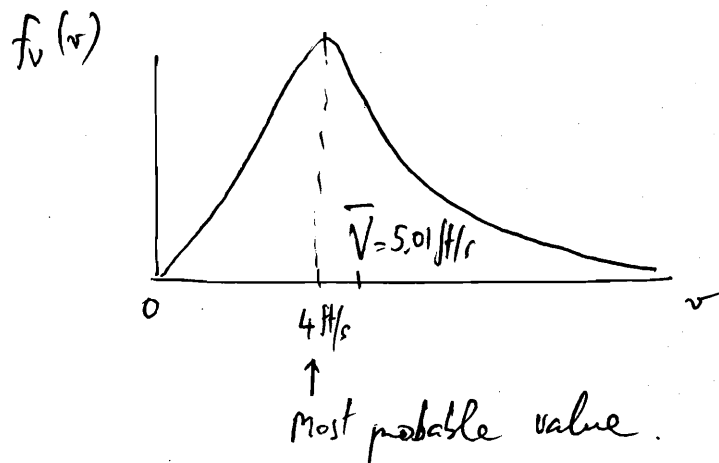
$$\frac{df_V}{dv} = 0 = \frac{1}{\sigma^2} e^{-\frac{v^2}{2\sigma^2}} - \frac{2v^2}{\sigma^2} e^{-\frac{v^2}{2\sigma^2}} = \underbrace{\left(1 - \frac{2v^2}{\sigma^2}\right)}_{\text{Non-zero}} \frac{e^{-\frac{v^2}{2\sigma^2}}}{\sigma^2}$$

$$\Rightarrow 1 - \frac{2v^2}{\sigma^2} = 0 \rightarrow v_{\text{most prob.}} = \frac{\sigma}{\sqrt{2}} = 4 \text{ ft/s}$$

$$\hookrightarrow \frac{v^2}{\sigma^2} = 1 \Rightarrow \frac{v}{\sigma} = 1$$

b) Mean value of speed

$$\bar{V} = \sigma \sqrt{\frac{\pi}{2}} = 4 \sqrt{\frac{\pi}{2}} = 5.01 \text{ ft/s}$$



$$c) P_2(v > 10 \text{ ft/s}) = 1 - P_2(v \leq 10 \text{ ft/s})$$

$$= 1 - F_v(10)$$

$$= 1 - \int_0^{10} dv f_v(v)$$

$$= 1 - \frac{1}{\sigma^2} \int_0^{10} v e^{-\frac{v^2}{2\sigma^2}} dv$$

$f_v(v) = 0 \quad v \leq 0$

Change of variable: $t = \frac{v^2}{2\sigma^2} \rightarrow dt = \frac{v dv}{\sigma^2}$

$$\begin{aligned}
 P_2(v > 10 \text{ ft/s}) &= 1 - \int_0^{10} dt e^{-t} \\
 &= 1 - \left[\frac{e^{-t}}{-1} \right]_{t=0}^{t=\frac{100}{2\sigma^2}} = 1 - \left[\frac{e^{-\frac{100}{2\sigma^2}} - 1}{-1} \right] \\
 &= e^{-\frac{100}{2\sigma^2}} = e^{-\frac{100}{32}} = 0.0439 \\
 &\quad \text{or } 4.39\%
 \end{aligned}$$

In general: Rayleigh distribution:

$$f_v(v) = \begin{cases} \frac{v}{\sigma^2} e^{-\frac{v^2}{2\sigma^2}} & v \geq 0 \\ 0 & v < 0 \end{cases}$$

$$\rightarrow F_v(v) = \int_0^v \frac{u}{\sigma^2} e^{-\frac{u^2}{2\sigma^2}} du = \int_0^{\frac{v^2}{2\sigma^2}} dt e^{-t}$$

$F_v(v) = 1 - e^{-\frac{v^2}{2\sigma^2}}$

$t = \frac{u^2}{2\sigma^2}$

3) Maxwell Distribution: 3D counterpart of Rayleigh distrib.

Gas molecules have speed components v_x, v_y, v_z that are Gaussian random variables with zero means and variance $\sigma_{v_x}^2 = \sigma_{v_y}^2 = \sigma_{v_z}^2 = \sigma^2 \rightarrow$ Total speed of the molecules is a random variable described by the Maxwell distribution:

$$f_v(v) = \begin{cases} \sqrt{\frac{2}{\pi}} \frac{v^2}{\sigma^3} e^{-\frac{v^2}{2\sigma^2}} & v \geq 0 \\ 0 & v < 0 \text{ (speed can't be negative)} \end{cases}$$

$$1) \quad \bar{v} = \int_0^\infty dv \, v f_v(v) = \sqrt{\frac{2}{\pi}} \int_0^\infty dv \, \frac{v^3}{\sigma^3} e^{-\frac{v^2}{2\sigma^2}} = \sqrt{\frac{2}{\pi}} \frac{1}{\sigma^3} \frac{\Gamma(2)}{2 \left(\frac{1}{2\sigma^2}\right)^4}$$

$$\boxed{\bar{v} = \sqrt{\frac{8}{\pi}} \sigma}$$

$$2) \quad \overline{v^2} = \int_0^\infty dv \, v^2 f_v(v) = \sqrt{\frac{2}{\pi}} \frac{1}{\sigma^3} \int_0^\infty dv \, v^4 e^{-\frac{v^2}{2\sigma^2}} = \sqrt{\frac{2}{\pi}} \frac{1}{\sigma^3} \frac{\Gamma(\frac{5}{2})}{2 \left(\frac{1}{2\sigma^2}\right)^5}$$

$$\boxed{\overline{v^2} = 3\sigma^2}$$

$$3) \quad \left[\sigma_v^2 = \overline{v^2} - \bar{v}^2 = 3\sigma^2 - \frac{8}{\pi} \sigma^2 = \left(3 - \frac{8}{\pi}\right) \sigma^2 = 0.453 \sigma^2 \right]$$

4) Chi-Square Distribution: $\chi^2 = Y_1^2 + Y_2^2 + \dots + Y_n^2$

Y_i = Gaussian random variables with zero means and unit variance $\rightarrow \chi^2$ follows the Chi-square distribution:

Conditional Probability Distribution and Density Function:

$$F(x|M) = P(\bar{X} \leq x | M) = \frac{P\{\bar{X} \leq x, M\}}{P(M)}$$

$\{\bar{X} \leq x, M\}$ is joint event of all outcomes $\bar{z}$ (x_i or 'x' in Greek alphabet) such that $\bar{X}(\bar{z}) \leq x$ and $\bar{z} \in M$

$F(x|M)$ is a valid probability distribution, if we check the three axioms:

1) $0 \leq F(x|M) \leq 1$ ✓ Since $P(M) \geq P\{\bar{X} \leq x, M\}$

2) $F(-\infty|M) = P(\bar{X} \leq -\infty | M) = 0$ ✓

$F(\infty|M) = P(\bar{X} \leq \infty | M) = 1$ ✓

3) $P(x_1 < \bar{X} \leq x_2 | M) = F(x_2|M) - F(x_1|M) \geq 0$ ✓

HW3 = due 3/13. 6.4; 7.1; 8.2; 9.3

(EX1 : HW1 → 3 & Ch1-2 on 3/25)

(6.4) a) $v_{\text{most probable}} = 4 \text{ ft/s}$ b) $\bar{V} = 5.01 \text{ ft/s}$ c) $P(v > 10 \text{ ft/s}) = 4.39\%$
Rayleigh distribution.

(7.1) a) $f_X(x) = \begin{cases} \frac{1}{2\pi\sqrt{1-x^2}} & -1 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$ b) $E[X] = 0$ c) $\overline{X^2} = \frac{1}{2}$
 $\sigma_X^2 = \frac{1}{2}$

d) $P(x > 0.5) = \frac{1}{3}$

(8.2) a) $f(x|M) = \frac{2}{\pi\sqrt{1-x^2}} \quad -1 \leq x \leq 1$

b) $E[X|M] = \frac{2}{\pi}$

(9.3) a) $f(v_0) = \begin{cases} \frac{1}{\sqrt{m}\sigma_{v_0}} e^{-\frac{(v_0 - \bar{v}_0)^2}{2\sigma_{v_0}^2}} & -B < v_0 < B \\ \Phi(-A) & v_0 = \pm B \\ 0 & \text{otherwise} \end{cases}$

b) $E[v_0] = 2.55$

HW3: (7.1)

 Θ : Random variable, uniformly distributed, b/w 0 & 2π $\bar{X}$: another random variable : $\bar{X} = \cos \Theta$ c) Find $f_{\bar{X}}(x)$:Recall: "Connection b/w two random variables": $\bar{X}$ & $\bar{Y}$:

$$f_{\bar{Y}}(y) = f_{\bar{X}}(x) \left| \frac{dx}{dy} \right|$$

In this problem: $\begin{matrix} X \rightarrow \Theta \\ Y \rightarrow \bar{X} \end{matrix} \rightarrow \boxed{f_{\bar{X}}(x) = f_{\Theta}(\theta) \left| \frac{d\theta}{dx} \right|}$

$$\begin{cases} X = \cos \Theta \rightarrow \frac{dx}{d\theta} = -\sin \Theta = -\sqrt{1 - \cos^2 \Theta} = -\sqrt{1 - x^2} \\ f_{\Theta}(\theta) = \begin{cases} \frac{1}{2\pi - 0} = \frac{1}{2\pi} & 0 \leq \theta \leq 2\pi \\ 0 & \text{otherwise} \end{cases} \end{cases}$$

$$\rightarrow f_{\bar{X}}(x) = \begin{cases} \frac{1}{2\pi} (\sqrt{1 - x^2})^{-1} & -1 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

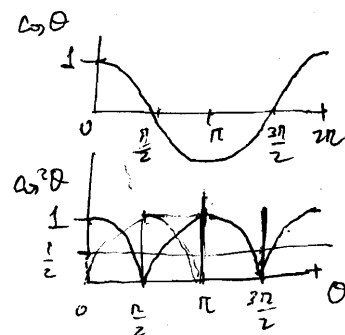
b) $\bar{X} = E[\bar{X}] = \begin{cases} \int_{-1}^1 dx \cdot x \cdot \frac{1}{2\pi} \frac{1}{\sqrt{1 - x^2}} & \text{: using standard definitions of the average using } f_{\bar{X}}(x) \\ \int_0^{2\pi} d\theta \cos \theta \cdot \frac{1}{2\pi} & \text{: using the connection with } \Theta \text{ and } f_{\Theta}(\theta) \end{cases}$

$= 0$

$$\rightarrow \bar{X} = E[\bar{X}] = 0$$

c) $\sigma_{\bar{X}}^2 = \overline{\bar{X}^2} - \bar{X}^2 = \overline{\bar{X}^2}$

$$\overline{\bar{X}^2} = E[\bar{X}^2] = \begin{cases} \int_{-1}^1 dx \cdot x^2 \cdot \frac{1}{2\pi} \frac{1}{\sqrt{1 - x^2}} \\ \int_0^{2\pi} d\theta \cos^2 \theta \cdot \frac{1}{2\pi} = \frac{1}{2} \end{cases}$$



$$\frac{1}{2\pi} \int_0^{2\pi} d\theta \cos^2 \theta = \text{integration by parts: } \int u dv = uv - \int v du$$

$$\begin{aligned} u &\equiv \cos \theta & \rightarrow du &= -\sin \theta d\theta \\ dv &\equiv \cos \theta d\theta & \rightarrow v &= \sin \theta \end{aligned}$$

$$\rightarrow \frac{1}{2\pi} \left[\underbrace{\cos \theta \sin \theta}_0 \Big|_0^{2\pi} + \int_0^{2\pi} d\theta \sin^2 \theta \right]$$

$$= \frac{1}{2\pi} \left[\int_0^{2\pi} d\theta (1 - \cos^2 \theta) \right] = \frac{1}{2\pi} \left[\int_0^{2\pi} d\theta - \int_0^{2\pi} d\theta \cos^2 \theta \right]$$

$$\Rightarrow \frac{2}{2\pi} \int_0^{2\pi} d\theta \cos^2 \theta = 1 \quad \rightarrow \quad \boxed{\frac{1}{2\pi} \int_0^{2\pi} d\theta \cos^2 \theta = \frac{1}{2}}$$

$$\Rightarrow E[X^2] = \overline{X^2} = \frac{1}{2} \quad \Rightarrow \quad \sigma_X^2 = \overline{X^2} = \frac{1}{2}$$

$$\begin{aligned} d) \quad P_2(X > 0.5) &= 1 - P_2(X \leq 0.5) = 1 - F_X(0.5) = 1 - \int_{-\infty}^{0.5} dx f_X(x) \\ &= 1 - \int_{-1}^{0.5} dx \frac{1}{2\pi} \frac{1}{\sqrt{1-x^2}} = \int_{0.5}^1 dx \frac{1}{2\pi} \frac{1}{\sqrt{1-x^2}} \end{aligned}$$

$$\boxed{\begin{array}{l} \text{In general: } \int \frac{dx}{\sqrt{1-x^2}} \stackrel{\downarrow}{=} \int d\theta \\ x \equiv \sin \theta \\ dx = \cos \theta d\theta \end{array}}$$

$$\begin{aligned} &= \frac{1}{2\pi} \int_{\sin^{-1} 0.5}^{\sin^{-1} 1} d\theta = \frac{1}{2\pi} \left[\frac{\pi}{2} - \frac{\pi}{6} \right] \\ &= \frac{1}{2\pi} \cdot \frac{4\pi}{12} = \boxed{\frac{1}{6}} \end{aligned}$$

(60)

8.2 c) For X in 7.1, $f(x|M)$? where M is the event of $0 \leq \theta \leq \frac{\pi}{2}$.

$$F(x | 0 \leq \theta \leq \frac{\pi}{2}) = P(X \leq x | 0 \leq \theta \leq \frac{\pi}{2}) = \frac{P(X \leq x, 0 \leq \theta \leq \frac{\pi}{2})}{P(0 \leq \theta \leq \frac{\pi}{2})}$$

$\downarrow$
 $X = \cos \theta$

$$\theta : \text{uniform dist. l/w } 0 \text{ \& } \pi \rightarrow P(0 \leq \theta \leq \frac{\pi}{2}) = \frac{1}{4}$$

$$\Rightarrow F(x | 0 \leq \theta \leq \frac{\pi}{2}) = 4 P(X \leq x, 0 \leq x \leq 1) = 4 F(x) \text{ if } 0 \leq x \leq 1$$

derivative w.r.t x $\rightarrow f(x|M) = 4 f_x(x) = \frac{4}{\pi} \frac{1}{\sqrt{1-x^2}} = \frac{2}{\pi \sqrt{1-x^2}} \text{ if } 0 \leq x \leq 1$

$$b) E[X|M] = E[\cos \theta | 0 \leq \theta \leq \frac{\pi}{2}] = \int_0^{\frac{\pi}{2}} d\theta \cos \theta \underbrace{\frac{4}{\pi}}_{f(\theta|M)}$$

use connection with θ
 $\rightarrow$ sin, cos integral

$$\hookrightarrow \text{As } f(x|M) = 4 f_x(x) = \frac{4}{\pi \sqrt{1-x^2}}$$

$$f(\theta|M) = 4 f_\theta(\theta) = \frac{4}{\pi}$$

$$= \frac{2}{\pi} \int_0^{\frac{\pi}{2}} d\theta \cos \theta = \frac{2}{\pi} [\sin \theta]_0^{\frac{\pi}{2}} = \frac{2}{\pi}$$

Alternatively:

$$E[X|M] = \int_0^1 dx \cdot x \cdot f(x|M) = \int_0^1 dx \cdot x \cdot \frac{2}{\pi \sqrt{1-x^2}} = \frac{2}{\pi} \int_0^1 dx \frac{x}{\sqrt{1-x^2}}$$

$$= \frac{2}{\pi} [-\cos \theta]_0^{\frac{\pi}{2}} = \frac{2}{\pi}$$

In general: $\int \frac{x dx}{\sqrt{1-x^2}} \stackrel{x = \cos \theta}{=} \int \sin \theta d\theta$

9.3

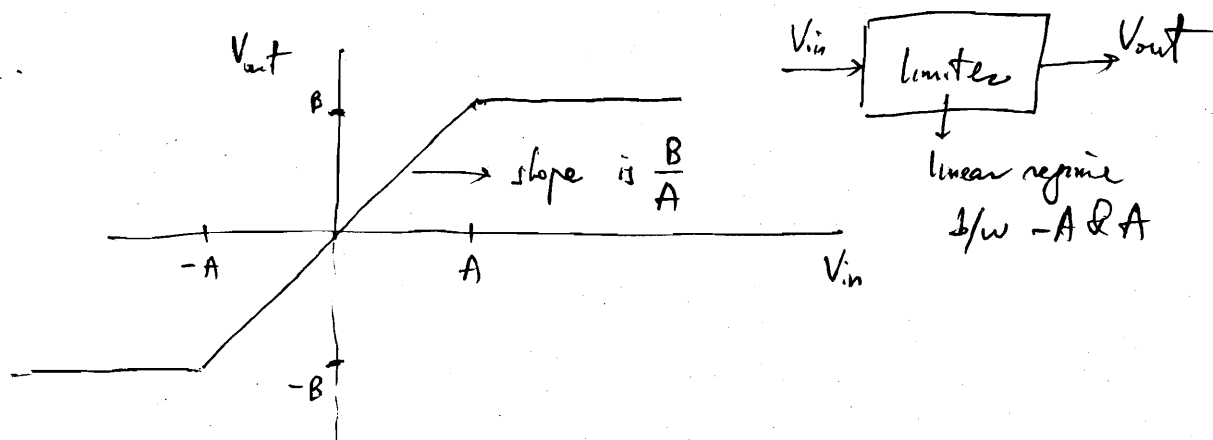
$$V_{out} = \begin{cases} -B \\ \left(\frac{B}{A} V_{in}\right) \\ B \end{cases}$$

$$V_{in} < -A$$

$$-A < V_{in} < A$$

$$V_{in} > A$$

limiter.



a) $V_{in} = \text{GRV } V$; mean $\bar{V}$, variance σ_V^2

write $f_{V_{out}} = f_{V_0}(v_0) = \begin{cases} \frac{1}{\sqrt{2\pi}\sigma_V} e^{-\frac{(v_0 - \bar{V})^2}{2\sigma_V^2}} & -B < v_0 < B \\ f_{V_{in}}(\pm A) & v_0 = \pm B \\ 0 & \text{otherwise} \end{cases}$

$$v_0 = +B \rightarrow v_{in} = A \rightarrow f_{V_{in}}(A)$$

$$v_0 = -B \rightarrow v_{in} = -A \rightarrow f_{V_{in}}(-A)$$

b) $E[v_0] = 2.55$